Mathematics. — "Weights of the most probable values of the Unknowns in the case of Direct Conditioned Observations". By Prof. M. J. VAN UVEN. (Communicated by Prof. JAN DE VRIES).

(Communicated at the meeting of October 30, 1926).

In adjusting direct conditioned observations, the reliableness of the "solutions", i. e. of the values obtained by adjustment, is indicated by their weights.

In the following paper we propose deriving for these weights an expression which gives them in a closed form as a function of the data, and which moreover holds good under an arbitrary number of conditions.

The quantities observed (*m* in number) will be designated by x, y, z, ...; the observed values may be $\xi, \eta, \zeta, ...,$ resp., their weights $g_{\xi}, g_{\chi}, g_{\zeta}, ...$

For simplification's sake we will assume that these m quantities x, y, z, \ldots must satisfy a priori μ linear conditions:

$$a_j x + \beta_j y + \gamma_j z + \ldots = \varkappa_j \qquad j = 1, 2, \ldots \mu, \quad \mu < m \quad . \quad (1)$$

We form the differences

$$\varkappa_{j} - (a_{j} \xi + \beta_{j} \eta + \gamma_{j} \zeta + \ldots) = \varkappa_{j} - [a_{j} \xi] = V_{j}, \quad j = 1, 2, \ldots \mu$$
 (2)

where [] denotes a summation over the variables x, y, z, \ldots Let X, Y, Z, \ldots be *m* provisionally assumed comparative values for x, y, z, \ldots , which, by agreement, satisfy the μ conditions, so that

$$[a_j X] = \varkappa_j \qquad j = 1, 2, \ldots \mu,$$

Then, by forming the differences

we have, according to (2):

$$[a_j v_x] = -V_j$$
 $j = 1, 2, \ldots \mu$ (4)

We next consider those comparative values X, Y, Z, \ldots as the most probable values for x, y, z, \ldots , for which $[g_{\xi} v_x v_x]$ takes the least value allowed by the conditions.

Then the equation $[g_{\xi} v_x dv_x] = 0$ must be dependent on the μ equations $[a_j dv_x] = 0$ $(j = 1, ..., \mu)$, whence

$$[g_{\xi} v_x dv_x] \equiv \sum_{j=1}^{\mu} \lambda_j [a_j dv_x];$$

here the λ_j are multiplicators to be determined; the summation sign Σ is used here (and will be used also in future) to designate a summation over the μ conditions. Thus the sum Σ runs always from 1 to μ .

Now we have, as a consequence of the above identity:

 $g_{\xi} v_x = \Sigma \lambda_j a_j$, $g_{\gamma} v_y = \Sigma \lambda_j \beta_j$, $g_{\zeta} v_z = \Sigma \lambda_j \gamma_j$,... *m* equations (5) and the μ equations already found:

hence altogether $m + \mu$ equations to determine the $m + \mu$ unknowns $v_x, v_y, v_z, \ldots, \lambda_1, \lambda_2, \ldots, \lambda_{\mu}$.

The usual solution runs as follows:

From (4) and (5) we derive the μ equations:

$$\sum_{j=1}^{\mu} \lambda_j \left[\frac{a_j \ a_k}{g_{\xi}} \right] = -V_k , \qquad k = 1, \ldots \mu$$

or, putting for the sake of abbreviation:

These μ equations determine the μ multiplicators λ_j .

Denoting the solution of these equations by $\overline{\lambda_j}$, and representing the values of v_x, v_y, v_z, \ldots , corresponding with them, according to (5), by u_x, u_y, u_z, \ldots , we have:

$$u_x = \frac{1}{g_{\xi}} \Sigma a_j \overline{\lambda}_j, \quad u_y = \frac{1}{g_{\gamma}} \Sigma \beta_j \overline{\lambda}_j, \quad u_z = \frac{1}{g_{\zeta}} \Sigma \gamma_j \overline{\lambda}_j, \ldots$$
 (8)

the most probable values of x, y, z, \ldots being

$$\overline{x} = \xi - u_x$$
, $\overline{y} = \eta - u_y$, $\overline{z} = \zeta - u_z$ (9)

and $[g_{\xi} u_x u_x]$ being the least value which $[g_{\xi} v_x v_x]$ may assume, so far as the conditions allow.

The mean (true) error σ of the unity of weight follows from

The sum $[g_{\xi} u_x u_x]$ may also be expressed directly into the data:

$$[g_{\xi} u_{x} u_{x}] = \sum_{k} \sum_{j} \left[g_{\xi} \cdot \frac{a_{k} \overline{\lambda}_{k}}{g_{\xi}} \cdot \frac{a_{j} \overline{\lambda}_{j}}{g_{\xi}} \right] = \sum_{k} \sum_{j} \left[\frac{a_{k} a_{j}}{g_{\xi}} \overline{\lambda}_{k} \overline{\lambda}_{j} \right] =$$

$$= \sum_{k} \sum_{j} \left[\frac{a_{k} a_{j}}{g_{\xi}} \right] \overline{\lambda}_{k} \overline{\lambda}_{j} = \sum_{k} \sum_{j} s_{kj} \overline{\lambda}_{k} \overline{\lambda}_{j}$$

(see (8) and (6)), or, since $\overline{\lambda_j}$, being the solution of (7), satisfies

$$\sum_{j} s_{kj} \overline{\lambda_{j}} = -V_{k},$$

$$[g_{\xi} u_{x} u_{x}] = -\sum_{k} V_{k} \overline{\lambda_{k}}.$$

Representing the determinant $|s_{jk}| = \begin{vmatrix} s_{11}, s_{12}, \dots s_{1\mu} \\ \vdots \\ s_{1\mu}, s_{2\mu}, \dots s_{\mu\mu} \end{vmatrix}$ by S, and the minor (algebraic complement) of the element s_{jk} by S_{jk} (whence $S = \sum_{i} s_{jk} S_{jk}$, $0 = \sum_{i} s_{jk} S_{jl} (l \neq k)$), we have

$$\overline{\lambda}_k = \frac{-\sum S_{jk} V_j}{S},$$

hence

$$[g_{\xi} u_{x} u_{x}] = \frac{\sum_{k} \sum_{j} S_{jk} V_{j} V_{k}}{S} = -\frac{1}{S} \begin{vmatrix} s_{11}, \dots, s_{1\mu}, V_{1} \\ \vdots \\ s_{1\mu}, \dots, s_{\mu\mu}, V_{\mu} \\ V_{1}, \dots, V_{\mu}, 0 \end{vmatrix},$$

or, putting for the sake of abbreviation:

$$\begin{vmatrix} s_{11}, \dots, s_{1\mu}, q_1 \\ \vdots & \vdots & \vdots \\ s_{1\mu}, \dots, s_{\mu\mu}, q_{\mu} \\ p_1, \dots, p_{\mu}, r \end{vmatrix} = \begin{vmatrix} s_{jk} q_j \\ p_k r \end{vmatrix}, \quad ... \quad .. \quad .. \quad (11)$$

$$[g_{\xi} u_{x} u_{x}] = -\frac{1}{S} \begin{vmatrix} s_{jk}, V_{j} \\ V_{k}, 0 \end{vmatrix}$$
 (12)

The uncertainty of the "solutions" \overline{x} , \overline{y} , \overline{z} , ... (see (9)) depends, in the last instance, on the uncertainty of the observed values ξ . η , ζ ,; these latter have the weights g_{ξ} , g_{γ} , g_{ζ} , ..., thus the mean errors:

$$\sigma_{\xi} = \frac{\sigma}{\sqrt{g_{\xi}}}, \quad \sigma_{\eta} = \frac{\sigma}{\sqrt{g_{\eta}}}, \quad \sigma_{\xi} = \frac{\sigma}{\sqrt{g_{\xi}}}, \dots$$

If we imagine that, in eventually repeating the series of observations, the observed values ξ . η , ζ ,..., undergo the variations $\Delta \xi$, $\Delta \eta$, $\Delta \zeta$,... respectively, then the mean square of $\Delta \xi$ (represented by $\overline{\Delta \xi^2}$) is nothing else than the square of the mean error of ξ , etc., in other words:

$$\overline{\Delta\xi^2} = \frac{\sigma^2}{g_{\xi}}, \quad \overline{\Delta\eta^2} = \frac{\sigma^2}{g_{\eta}}, \quad \overline{\Delta\zeta^2} = \frac{\sigma^2}{g_{\xi}}, \dots \dots \dots \dots (13)$$

the mutual independency of the observed values ξ , η , ζ ,... having as a consequence that the mean products, two by two, of the variations $\Delta \xi$, $\Delta \eta$, $\Delta \zeta$,... are zero, so that

$$\overline{\Delta \xi \cdot \Delta \eta} = 0, \quad \overline{\Delta \xi \cdot \Delta \zeta} = 0, \dots, \overline{\Delta \eta \cdot \Delta \zeta} = 0 \text{ etc.} \quad . \quad (14)$$

In consequence of the variations $\Delta \xi$, $\Delta \eta$, $\Delta \zeta$, . . ., the V_j undergo (according to (2)) the variations

The variations $\Delta \lambda_k$ of $\overline{\lambda_k}$, deriving from these ΔV_j , satisfy (see (7))

$$\Sigma s_{kj} \bigtriangleup \lambda_j = -\bigtriangleup V_k \, , \qquad k = 1, \ldots \mu \quad . \quad . \quad . \quad (15)$$

Further, from

$$u_x = \frac{1}{g_{\xi}} \sum_j a_j \overline{\lambda_j} = \xi - \overline{x} : \ldots \quad (8), \quad (9)$$

follows

We now multiply the μ equations (15) and the equation (16) by $\Delta \xi$. Then, taking the mean value, we obtain (by (13) and (14))

$$\sum_{j} S_{kj} \overline{\Delta \lambda_{j} \cdot \Delta \xi} = -\overline{\Delta V_{k} \cdot \Delta \xi} = + (a_{k} \overline{\Delta \xi^{2}} + \beta_{k} \overline{\Delta \xi \cdot \Delta \eta} + \gamma_{k} \overline{\Delta \xi \cdot \Delta \zeta} + ...)$$

$$= \frac{a_{k}}{g_{\xi}} \sigma^{2} \qquad k = 1, \dots, \mu$$

$$\sum \frac{a_{j}}{g_{\xi}} \overline{\Delta \lambda_{j} \cdot \Delta \xi} = \overline{\Delta \xi^{2}} - \overline{\Delta \overline{x} \cdot \Delta \xi} = \frac{\sigma^{2}}{g_{\xi}} - \overline{\Delta \overline{x} \cdot \Delta \xi}.$$

By eliminating the μ variables $\overline{\bigtriangleup \lambda_j}$. $\bigtriangleup \xi$ from these $\mu + 1$ equations and using the notation (11), we arrive at

$$\frac{s_{kj}}{g_{\xi}}, \frac{\frac{a_k}{g_{\xi}}}{\frac{a_j}{g_{\xi}}} - \overline{\bigtriangleup x} \cdot \bigtriangleup \xi} = 0,$$

or

$$\left| \begin{array}{c} s_{kj}, \ \displaystyle rac{a_k}{g_\xi} \ \sigma^2 \\ \displaystyle rac{a_j}{g_\xi}, \ \displaystyle rac{\sigma^2}{g_\xi} \end{array}
ight| = \left| \begin{array}{c} s_{kj}, \ 0 \\ \displaystyle rac{a_j}{g_\xi}, \ \displaystyle ar{\Delta \, \overline{x} \, . \, igsires \xi} \end{array}
ight| = 0,$$

or

$$\frac{\sigma^2}{g_{\xi}} \begin{vmatrix} s_{kj}, & \frac{a_k}{g_{\xi}} \\ a_j, & 1 \end{vmatrix} - S.\overline{\Delta x}.\Delta \xi = 0,$$

l

whence

In the same manner we obtain, by multiplying the μ equations (15) and the equation (16) by $\Delta \eta$, and then taking the mean, $\sum_{j} s_{kj} \overline{\Delta \lambda_{j} . \Delta \eta} = -\overline{\Delta V_{k} . \Delta \eta} = + (a_{k} \overline{\Delta \eta . \Delta \xi} + \beta_{k} \overline{\Delta \eta^{2}} + \gamma_{k} \overline{\Delta \eta . \Delta \zeta} + ...)$ $= \frac{\beta_{k}}{g_{\tau}} \sigma^{2} \qquad k = 1, ... \mu$ $\sum_{j} \frac{a_{j}}{g_{\xi}} \overline{\Delta \lambda_{j} . \Delta \eta} = \overline{\Delta \xi . \Delta \eta} - \overline{\Delta x} . \Delta \eta = -\overline{\Delta x} . \overline{\Delta \eta},$ From these $\mu + 1$ equations we derive by eliminating the μ variables $\overline{\bigtriangleup \lambda_j \, . \, \bigtriangleup \eta}$:

$$\begin{vmatrix} s_{kj}, & \frac{\beta_k}{g_{\zeta}} \sigma^2 \\ \frac{a_j}{g_{\zeta}}, & 0 - \overline{\bigtriangleup \, \overline{x} \, . \, \bigtriangleup \, \eta} \end{vmatrix} = 0,$$

or

$$\begin{vmatrix} s_{kj}, & \frac{\beta_k}{g_{\tau_i}} \sigma^2 \\ \frac{a_j}{g_{\xi}}, & 0 \end{vmatrix} - \begin{vmatrix} s_{kj}, & 0 \\ \frac{a_j}{g_{\xi}}, & \overline{\Delta \, \overline{x} \, . \, \Delta \, \eta} \end{vmatrix} = 0,$$

or

$$\frac{\sigma^2}{g_{\xi}} \begin{vmatrix} s_{kj}, & \frac{\beta_k}{g_{\tau}} \\ a_j, & 0 \end{vmatrix} - S \cdot \overline{\Delta \, \overline{x} \, \Delta \, \eta} = 0,$$

whence

$$\overline{\bigtriangleup \overline{x} \, . \, \bigtriangleup \eta} = \frac{1}{S} \, . \, \frac{\sigma^2}{g_{\xi}} \, . \, \left| \begin{array}{cccc} s_{kj}, & \frac{\beta_k}{g_{\eta}} \\ a_j, & 0 \end{array} \right| \quad . \quad . \quad . \quad . \quad (18)$$

Likewise we find

$$\overline{\bigtriangleup \overline{x} \, . \, \bigtriangleup \zeta} = \frac{1}{S} \, . \, \frac{\sigma^2}{g_{\xi}} \, . \, \left| \begin{array}{ccc} s_{kj}, & \frac{\gamma_k}{g_{\xi}} \\ a_j & , & 0 \end{array} \right| \, . \, . \, . \, . \, (18 \text{ bis})$$

etc.

We next multiply (15) by $\triangle x$ and take the mean; so we obtain (replacing the index k by i):

$$\sum_{j} s_{ij} \overline{\Delta \lambda_{j} \cdot \Delta \overline{x}} = -\overline{\Delta V_{i} \cdot \Delta \overline{x}} = + (a_{i} \overline{\Delta \overline{x} \cdot \Delta \xi} + \beta_{i} \overline{\Delta \overline{x} \cdot \Delta \eta} + \gamma_{i} \overline{\Delta \overline{x} \cdot \Delta \zeta} + \dots), i = 1, \dots, \mu$$

or, on account (17), (18), (18 bis) etc.

$$\begin{split} \sum_{j} s_{ij} \overline{\Delta \lambda_{j} \cdot \Delta \overline{x}} &= \frac{1}{S} \cdot \frac{\sigma^{2}}{g_{\xi}} \cdot \left\{ a_{i} \left| \begin{array}{c} s_{kj} \cdot \frac{a_{k}}{g_{\xi}} \\ a_{j} & \cdot 1 \end{array} \right| + \beta_{i} \left| \begin{array}{c} s_{kj} \cdot \frac{\beta_{k}}{g_{\tau_{i}}} \\ a_{j} & \cdot 0 \end{array} \right| + \gamma_{i} \left| \begin{array}{c} s_{kj} \cdot \frac{\gamma_{k}}{g_{\xi}} \\ a_{j} & \cdot 0 \end{array} \right| + \cdots \right\} \\ &= \frac{1}{S} \cdot \frac{\sigma^{2}}{g_{\xi}} \cdot \left| \begin{array}{c} s_{kj} \cdot \left[\frac{a_{k} \cdot a_{i}}{g_{\xi}} \right] \\ a_{j} & \cdot a_{i} \end{array} \right| = \frac{1}{S} \cdot \frac{\sigma^{2}}{g_{\xi}} \cdot \left| \begin{array}{c} s_{kj} \cdot s_{ki} \\ a_{j} & \cdot a_{i} \end{array} \right| \cdot i = 1, \dots, \mu \end{split}$$

or

$$\sum_{j} s_{ij} \overline{\Delta \lambda_{j} . \Delta \overline{x}} = \frac{1}{S} \cdot \frac{\sigma^{2}}{g_{\xi}} \begin{vmatrix} s_{11}, \ldots, s_{1i}, \ldots, s_{1\mu} , s_{1i} \\ \vdots \\ s_{\mu 1}, \ldots, s_{\mu i}, \ldots, s_{\mu \mu} , s_{\mu i} \\ a_{1}, \ldots, a_{i} , \ldots, a_{\mu} , a_{i} \end{vmatrix} \equiv 0 \qquad i = 1, \ldots, \mu$$

The determinant $|s_{ij}| = S$ of the μ homogeneous equations

$$\sum_{j} s_{ij} \overline{\bigtriangleup \lambda_j . \bigtriangleup x} = 0 \qquad i = 1, \ldots \mu$$

which we obtain in this way, being different from zero, we have separately

$$\overline{\bigtriangleup \lambda_j \, . \, \bigtriangleup \overline{x}} = 0. \qquad j = 1, \ldots \mu \quad . \quad . \quad . \quad . \quad . \quad (19)$$

Multiplying at last the equation (16) by $\triangle \overline{x}$ and then taking the mean, we find

$$\overline{\bigtriangleup u_x \,.\, \bigtriangleup \overline{x}} = \Sigma \, \frac{a_j}{g_{\xi}} \, \overline{\bigtriangleup \lambda_j \,.\, \bigtriangleup \overline{x}} = \overline{\bigtriangleup \, \xi \,.\, \bigtriangleup \, \overline{x}} - \overline{\bigtriangleup \, \overline{x}^2}$$

thus, by (19),

$$\overline{\bigtriangleup u_x \, . \, \bigtriangleup \overline{x}} = \overline{\bigtriangleup \xi \, . \, \bigtriangleup \overline{x}} - \overline{\bigtriangleup \overline{x}^2} = 0 \quad . \quad . \quad . \quad . \quad (20)$$

or, on account of (17),

$$\overline{\bigtriangleup x^2} = \overline{\bigtriangleup \xi . \bigtriangleup x} = \frac{1}{S} \cdot \frac{\sigma^2}{g_{\xi}} \begin{vmatrix} s_{kj} & \frac{a_k}{g_{\xi}} \\ a_j & 1 \end{vmatrix} \cdot \ldots \cdot \cdot \cdot (21)$$

Now, $\overline{\Delta \overline{x}^2}$ being nothing but the square of the mean error $\sigma_{\overline{x}}$ of \overline{x} , we have

$$\sigma_{\overline{x}}^{2} = \frac{1}{S} \cdot \frac{\sigma^{2}}{g_{\xi}} \cdot \begin{vmatrix} s_{kj} , \frac{a_{k}}{g_{\xi}} \\ a_{j} , 1 \end{vmatrix} \quad \dots \quad \dots \quad (22)$$

We shall moreover put this latter determinant into another form. In

$$\begin{vmatrix} s_{kj} &, \frac{a_k}{g_{\xi}} \\ a_j &, 1 \end{vmatrix} = \begin{vmatrix} s_{11}, \dots, s_{ij}, \dots, s_{1\mu}, a_1 : g_{\xi} \\ \vdots & \vdots & \vdots \\ s_{k1}, \dots, s_{kj}, \dots, s_{k\mu}, a_k : g_{\xi} \\ \vdots & \vdots & \vdots \\ s_{\mu 1}, \dots, s_{\mu j}, \dots, s_{\mu \mu}, a_{\mu} : g_{\xi} \\ a_1, \dots, a_j, \dots, a_{\mu}, 1 \end{vmatrix}$$

we take from the 1^{st} column a_1 times the last column

Then we retain for the element of the j^{th} column and the k^{th} row:

$$s_{kj} - \frac{a_j a_k}{g_{\xi}} = \left\lfloor \frac{a_k a_j}{g_{\xi}} \right\rfloor - \frac{a_k a_j}{g_{\xi}} = \frac{\beta_k \beta_j}{g_{\chi}} + \frac{\gamma_k \gamma_j}{g_{\chi}} + \ldots = a_{kj} \quad (23)$$

the j^{th} element of the last row becoming zero $(j = 1, \dots, \mu)$. Representing the determinant $|a_{kj}|$ by A, we obtain

$$\begin{vmatrix} s_{kj}, \frac{a_k}{g_{\xi}} \\ a_j, 1 \end{vmatrix} = \begin{vmatrix} a_{kj}, \frac{a_k}{g_{\xi}} \\ 0, 1 \end{vmatrix} = |a_{kj}| = A$$

So we find

$$\sigma_{\overline{x}}^2 = \frac{A}{S} \cdot \frac{\sigma^2}{g_{\xi}}$$

and for the weight $g_{\overline{x}}$ of \overline{x} :

$$g_{\overline{x}} = \frac{S}{A} \cdot g_{\xi}$$

Putting in a similar manner:

etc.,

and further:

$$|b_{kj}| = B$$
 , $|c_{kj}| = C$, etc.,

we arrive at last at:

$$\sigma_{\overline{x}}^2 = \frac{A}{S} \cdot \frac{\sigma^2}{g_{\xi}} \quad , \quad \sigma_{\overline{y}}^2 = \frac{B}{S} \cdot \frac{\sigma^2}{g_{\eta}} \quad , \quad \sigma_{\overline{z}}^2 = \frac{C}{S} \cdot \frac{\sigma^2}{g_{\xi}} , \quad \text{etc.,} \quad . \quad (24)$$

or

$$g_{\overline{x}} = \frac{S}{A} \cdot g_{\xi}$$
, $g_{\overline{y}} = \frac{S}{B} \cdot g_{\chi}$, $g_{\overline{z}} = \frac{S}{C} \cdot g_{\xi}$, etc. . . (25)

Thus we have expressed the weights of the solutions in a closed form into the data g_{ξ} , g_{η} , g_{ζ} , ...; a_j , β_j , γ_j , ... $(j = 1, ..., \mu)$ If there is given, in particular, but one condition

 $\alpha x + \beta y + \gamma z + \ldots = \varkappa$

so that $\mu = 1$, then we have

$$S = \left[\frac{a a}{g_{\xi}}\right], \ A = \left[\frac{a a}{g_{\xi}}\right] - \frac{a a}{g_{\xi}} = \frac{\beta \beta}{g_{\eta}} + \frac{\gamma \gamma}{g_{\xi}} + \dots,$$
$$B = \left[\frac{a a}{g_{\xi}}\right] - \frac{\beta \beta}{g_{\eta}} = \frac{a a}{g_{\xi}} + \frac{\gamma \gamma}{g_{\xi}} + \dots, \ C = \dots, \text{ etc.},$$

thus

$$\sigma_{\overline{x}}^{2} = \frac{\sigma^{2}}{g_{\xi}} \cdot \frac{\frac{\beta \beta}{g_{\chi}} + \frac{\gamma \gamma}{g_{\xi}} + \dots}{\frac{a a}{g_{\xi}} + \frac{\beta \beta}{g_{\chi}} + \frac{\gamma \gamma}{g_{\xi}} + \dots}, \quad \sigma_{\overline{y}}^{2} = \frac{\sigma^{2}}{g_{\chi}} \cdot \frac{\frac{a a}{g_{\xi}} + \frac{\gamma \gamma}{g_{\xi}} + \dots}{\frac{a a}{g_{\xi}} + \frac{\beta \beta}{g_{\chi}} + \frac{\gamma \gamma}{g_{\xi}} + \dots}, \quad \sigma_{\overline{z}}^{2} = \dots, \text{ etc.}$$
(24 bis)

The determinants S, A, B, C, ... may yet be put into another form. Of the *m* variables x, y, z, \ldots , represented by $g_{\xi}, g_{\chi}, g_{\zeta}, \ldots$ and by $a_j, \beta_j, \gamma_j, \ldots$, we consider the combinations μ by μ . Be $x, z, \ldots t$ $(g_{\xi}, g_{\zeta}, \dots, g_{\tau}; a_{j}, \gamma_{j}, \dots, \theta)$ such a combination. Then we may form determinants such as

$$\triangle (x, z, \ldots, t) = \begin{vmatrix} a_1 & , & \gamma_1 & \ldots & \theta_1 \\ a_2 & , & \gamma_2 & \ldots & \theta_2 \\ \vdots \\ a_{\mu} & , & \gamma_{\mu} & \ldots & \theta_{\mu} \end{vmatrix}$$

According to the theory of determinants, we may write S as follows:

$$S = |s_{kj}| = \left| \left[\frac{a_k a_j}{g_{\xi}} \right] \right| = \Sigma \frac{\Delta^2 (x, z, \dots, t)}{g_{\xi} \cdot g_{\xi} \cdots g_{\tau}} \dots \dots \dots (27)$$

where the sum Σ runs over all the $C_{\mu}(m) = \frac{m!}{\mu! (m-\mu)!}$ combinations μ by μ of the *m* elements x, y, z, \ldots

In the same manner we may also reduce the determinant A.

Here all the *m* elements x, y, z, \ldots appear, except the element *x*. With these m-1 elements y, z, \ldots we can build determinants of the form:

wherin the elements a_i are always missing.

So we have

$$A = |a_{kj}| = \left| \left[\frac{a_k a_j}{g_{\xi}} \right] - \frac{a_k a_j}{g_{\xi}} \right| = \Sigma_x \frac{\Delta_x^2(y, z, \dots, t)}{g_{\tau_i} \cdot g_{\xi} \cdot \dots \cdot g_{\tau}}, \quad . \quad (29)$$

where the sum Σ_x extends over all the C_{μ} $(m-1) = \frac{(m-1)!}{\mu! (m-\mu-1)!}$ combinations μ by μ of the m-1 elements y, z, \ldots (x absent). Similarly:

$$B = |b_{kj}| = \Sigma_y \frac{\triangle_y^2(x, z, \dots, t)}{g_{\xi} \cdot g_{\zeta} \cdot \dots \cdot g_{\tau}} \dots \dots \dots (29 \text{ bis})$$

$$C = |c_{kj}| = \Sigma_z \frac{\Delta_z^2(x, y, \dots, t)}{g_{\xi} \cdot g_{\tau} \cdot \dots \cdot g_{\tau}} \dots \dots \dots \dots (29 \text{ ter})$$

etc.

So the determinants S, A, B, C,... are written as sums of positive terms. The terms of the sum A also appear in the sum S, which contains besides those terms depending upon the element $x(g_{\xi}, a_j)$. Thus the value A is certainly less than that of S. Likewise B, C,... are altogether less than S.

The fractions $\frac{A}{S}$, $\frac{B}{S}$, $\frac{C}{S}$,... are therefore altogether less than unity, so that the mean errors $\sigma_{\overline{x}}$, $\sigma_{\overline{y}}$, $\sigma_{\overline{z}}$,... of the solutions are less than the mean errors of the respective observations, viz:

$$\sigma_{\xi} = \frac{\sigma}{\sqrt{g_{\xi}}}$$
, $\sigma_{\eta} = \frac{\sigma}{\sqrt{g_{\eta}}}$, $\sigma_{\xi} = \frac{\sigma}{\sqrt{g_{\xi}}}$,....

thus:

$$\sigma_{\overline{x}} < \sigma_{\xi}$$
, $\sigma_{\overline{y}} < \sigma_{\gamma}$, $\sigma_{\overline{z}} < \sigma_{\zeta}$, (30)

If all the terms of the sum S(27) are nearly equal, then the sums A and S are almost proportional to the numbers of their terms, thus as $C_{\mu}(m-1)$ to $C_{\mu}(m)$, that is: as $m-\mu$ to m. In this case the mean errors of the solutions are about $\sqrt{\frac{m-\mu}{m}}$ times as large as the mean errors of the corresponding observations. The larger μ is in comparison to m, the more the solution surpasses the observation in accuracy.